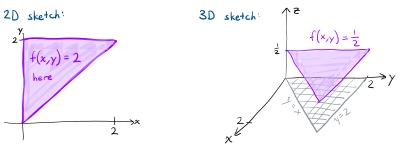
1. Let *X* and *Y* have joint density  $f(x, y) = \frac{1}{2}$  for  $0 \le x \le y \le 2$ .

(a) Sketch the joint density of *X* and *Y*.

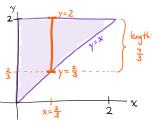


(b) What is the marginal density of *X*?

$$f_{\chi}(x) = \int_{x}^{z} f(x, y) \, dy = \int_{\chi}^{z} \frac{1}{2} \, dy = \frac{y}{2} \Big|_{y=x}^{y=2} = \frac{2-x}{2} \quad \text{for } O \le x \le 2$$

(c) Suppose you know that  $X = \frac{2}{3}$ . What does  $f\left(\frac{2}{3}, y\right)$  tell you about the density of *Y*, given that  $X = \frac{2}{3}$ ?

If  $x = \frac{2}{3}$ , then  $f(\frac{2}{3}, \gamma) = \frac{1}{2}$  for  $\frac{2}{3} \le \gamma \le 2$ . Since  $f(\frac{2}{3}, \gamma)$  is constant, the density of  $Y \mid X = \frac{2}{3}$  is constant for  $\frac{2}{3} \le \gamma \le 2$ , which is on interval of length  $\frac{4}{3}$ . Thus,  $f_{Y|X}(\gamma|\frac{2}{3}) = \frac{3}{4}$  for  $\frac{2}{3} \le \gamma \le 2$ . That is,  $Y \sim U_{ni}f[\frac{2}{3}, 2]$ .



(d) Suppose you know that  $X = x_0$ . What is then the density of *Y*?

In this case,  $Y \sim U_{nif}[x_0, 2]$ , so  $f_{Y|X}(y|x_0) = \frac{1}{2-x_0}$  for  $x_0 \leq y \leq 2$ .

(e) In part (d), you found the conditional density  $f_{Y|X}(y \mid x_0)$ . How does this relate to the joint density f(x, y) and the marginal density  $f_X(x)$ ?

Observe that 
$$f_{y|x}(y|x_{o}) = \frac{1}{2-x_{o}} = \frac{\frac{1}{2}}{\frac{2-x_{o}}{2}} = \frac{f(x_{o}, y)}{f_{x}(x_{o})}$$

(f) If  $X = \frac{2}{3'}$  then what is the probability that  $Y \le 1$ ?

$$P(Y = 1 \mid X = \frac{2}{3}) = \int_{\frac{2}{3}}^{1} f_{Y|X}(y|\frac{2}{3}) dy = \int_{\frac{2}{3}}^{1} \frac{3}{4} dy = \frac{3}{4} \cdot \frac{1}{3} = \frac{1}{4}$$

$$(1)$$

$$Use the conditional density of Y given  $X = \frac{2}{3}$$$

(g) What is the expected value of *Y* given that  $X = x_0$ ?

mean of Unif[xo, 2]

$$E(Y | X = x_{o}) = \int_{x_{o}}^{z} \gamma \cdot f_{Y|X}(y | x_{o}) d\gamma = \int_{x_{o}}^{z} \gamma \cdot \frac{1}{2 - x_{o}} d\gamma = \frac{1}{2(2 - x_{o})} \gamma^{2} \Big|_{y = x_{o}}^{y = 2} = \frac{4 - x_{o}^{2}}{2(2 - x_{o})} = \frac{2 + x_{o}}{2}$$

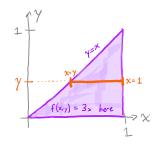
- 2. The joint pdf of *X* and *Y* is f(x, y) = 3x, for  $0 \le y \le x \le 1$ .
- (a) What is the conditional distribution of *X* given Y = y?

$$f_{\gamma}(\gamma) = \int_{\gamma}^{1} 3x \, dx = \frac{3}{2} x^{2} \Big|_{x=\gamma}^{x=1} = \frac{3}{2} (1-\gamma^{2}) \quad \text{for} \quad 0 \le \gamma \le 1$$

$$f_{\chi|\gamma}(\chi|\gamma) = \frac{f(\chi,\gamma)}{f_{\gamma}(\gamma)} = \frac{3\chi}{\frac{3}{2}(1-\gamma^{2})} = \frac{2\chi}{1-\gamma^{2}} \quad 0 \le \gamma \le \chi \le 1$$

$$f_{\chi|\gamma}(\chi|\gamma) = \frac{f(\chi,\gamma)}{f_{\gamma}(\gamma)} = \frac{3\chi}{\frac{3}{2}(1-\gamma^{2})} = \frac{2\chi}{1-\gamma^{2}} \quad 0 \le \gamma \le \chi \le 1$$

$$f_{\chi|\gamma}(\chi|\gamma) = \frac{f(\chi,\gamma)}{f_{\chi}(\chi)} = \frac{3\chi}{\frac{3}{2}(1-\gamma^{2})} = \frac{2\chi}{1-\gamma^{2}} \quad 0 \le \gamma \le \chi \le 1$$



(b) What is E(X | Y = y)?

$$E(X | Y = \gamma) = \int_{\gamma}^{1} \chi \cdot \frac{2x}{1 - \gamma^{2}} dx = \frac{2}{1 - \gamma^{2}} \int_{\gamma}^{1} \chi^{2} dx = \frac{2}{1 - \gamma^{2}} \cdot \frac{x^{3}}{3} \Big|_{x = \gamma}^{x = 1} = \frac{2}{1 - \gamma^{2}} \left(\frac{1}{3} - \frac{\gamma^{3}}{3}\right) = \frac{2(1 - \gamma^{3})}{3(1 - \gamma^{2})}$$

(c) What is Var(X | Y = y)?

$$E(X^{2} | Y = \gamma) = \int_{\gamma}^{1} \chi^{2} \cdot \frac{2\chi}{1 - \gamma^{2}} d\chi = \frac{1}{1 - \gamma^{2}} \cdot \frac{\chi^{q}}{2} \Big|_{x=\gamma}^{x=1} = \frac{1 - \gamma^{q}}{2(1 - \gamma^{2})} = \frac{1}{2} (1 + \gamma^{2})$$

$$Var(X^{2} | Y = \gamma) = \frac{1}{2} (1 + \gamma^{2}) - \left(\frac{2(1 - \gamma^{3})}{3(1 - \gamma^{2})}\right)^{2} = \frac{(\gamma - 1)^{2}(\gamma^{2} + 4\gamma + 1)}{18(1 + \gamma)^{2}}$$

3. For continuous random variables *X* and *Y*, show that E(E(X | Y)) = E(X).

$$E(X | Y=\gamma) = \int_{\infty}^{\infty} x \cdot f_{x|Y}(x|\gamma) dx = \int_{-\infty}^{\infty} x \cdot \frac{f(x,\gamma)}{f_{Y}(\gamma)} dx, \text{ which is a function of } y$$

$$E(E(X | Y)) = \int_{\infty}^{\infty} E(X | Y=\gamma) f_{Y}(\gamma) d\gamma = \int_{\infty}^{\infty} \int_{-\infty}^{\infty} x \cdot \frac{f(x,\gamma)}{f_{Y}(\gamma)} dx f_{Y}(\gamma) d\gamma$$

$$= \int_{\infty}^{\infty} \int_{-\infty}^{\infty} x f(x,\gamma) d\gamma dx = \int_{\infty}^{\infty} x \cdot \frac{f(x,\gamma)}{f_{X}(x)} dx$$

$$= \int_{-\infty}^{\infty} x \cdot f_{x}(x) dx = E(X)$$

4. The number of eggs *N* found in a nest of a certain species of turtle has a Poisson distribution with mean  $\lambda$ . Each egg has a probability *p* of being viable, and this event is independent from egg to egg. Find the mean and variance of the number of viable eggs per nest.

rvs: 
$$N \sim Poisson(\lambda)$$
,  $X = number of viable eggs ~ Bin(N, p)$   
mean:  $E(X) = E(E(X|N)) = E(Np) = p E(N) = p \lambda$   
 $E(X|N) = Np^{p}$   
variance:  $Var(X) = Var(E(X|N)) + E(Var(X|N)) = Var(Np) + E(Np(1-p))$   
 $= p^{2} Var(N) + p(1-p) E(N) = p^{2}\lambda + p(1-p)\lambda = p^{2}\lambda - p^{2}\lambda + p\lambda = p\lambda$ 

**BONUS:** If *X* and *Y* are independent binomial random variables with identical parameters *n* and *p*, calculate the conditional expected value of *X* given that X + Y = m.

First, compute the conditional pmf of X given that 
$$X + Y = m$$
.  

$$P(X = k \mid X + Y = m) = \frac{P(X = k \text{ and } X + Y = m)}{P(X + Y = m)} = \frac{P(X = k) P(Y = m - k)}{P(X + Y = m)}$$

$$= \frac{\binom{n}{k} p^{k} (1 - p)^{n-k} \cdot \binom{n}{m-k} p^{m-k} (1 - p)^{n-m+k}}{\binom{2n}{m} p^{m} (1 - p)^{2n-m}} \qquad \left[ \begin{array}{c} \text{For the denominator, note that}}{X + Y - Bin(2n, p)} \right]$$

$$= \frac{\binom{n}{k} \binom{n}{m-k}}{\binom{2n}{m}}$$

This is a hypergeometric probability: the probability of k successes in a sample of size m from a population with n successes and n failures.

So the conditional distribution of X, given that X + Y = m, is hypergeometric, and its mean is  $E(X | X + Y = m) = \frac{m}{2}$ .