- 1. Suppose that emails arrive in your inbox according to a Poisson process with rate 2 emails per hour. Then the time between successive emails is an exponential random variable with mean 30 minutes.
- (a) What is the probability that an email arrives in the next 20 minutes?

(b) What is the probability that you don't receive any emails in the next hour?

We can use the exponential distribution:

$$X \sim Exp(x=2)$$
, so $P(X > 1) = \int_{1}^{\infty} 2e^{-2x} dx = -e^{-2x} \Big|_{1}^{\infty} = e^{-2} \approx 0.135$
 $E(X) = \frac{1}{x} = \frac{1}{2}$ hour = 30 minutes

Alternatively, use the Poisson distribution: $Y \sim \text{Poisson}(\mu=2)$, so $P(Y=0) = e^{-2} \frac{2^0}{D!} = e^{-2}$ $e^{-\pi} \frac{\mu^{\pi}}{\pi!}$

L-Xx

(c) What is the standard deviation of the time until the next email?

$$\sigma_x = E(X) = \frac{1}{2}$$
 hour = 30 minutes

2. Let
$$X \sim \text{Exp}(\lambda)$$
 and $0 < a < b$.
(a) What is $P(X \ge a)$?
Use the pdf: $P(X \ge a) = \int_{a}^{\infty} \lambda e^{-\lambda x} dx = -e^{-\lambda x} \Big|_{a}^{\infty} = O - (-e^{-\lambda a}) = \boxed{e^{-\lambda a}}$
exponential tail probability
Or use the cdf: $P(X \le a) = \int_{a}^{a} \lambda e^{-\lambda x} dx = -e^{-\lambda x} \Big|_{a}^{a} = -e^{-\lambda a} + e^{a} = 1 - e^{-\lambda a}$
Thus, $P(X \ge a) = 1 - (1 - e^{-\lambda a}) = \boxed{e^{-\lambda a}}$

(b) Show that
$$P(X > b | X > a) = P(X > b - a)$$
.
 $P(X > b | X > a) = \frac{P(X > b - a)}{P(X > a)} = \frac{P(X > b)}{P(X > a)} = \frac{P(X > b)}{P(X > a)} = \frac{e^{-\lambda b}}{e^{-\lambda a}}$

definition of
 $e^{-\lambda (b-a)} = P(X > b-a)$

(c) What other distribution satisfies the equality in (b)?

(d) The property in (b) is special, in the sense that it doesn't hold for most random variables. For example, if $U \sim \text{Unif}[0,10]$, show that $P(U > 4 | U > 3) \neq P(U > 1)$.

$$P(U>4 | U>3) = \frac{P(U>4)}{P(U>3)} = \frac{\frac{6}{10}}{\frac{7}{10}} = \frac{6}{7} \neq P(U>1) = \frac{9}{10}$$
$$P(U>b | U>a) \neq P(U>b-a)$$

3. What is the moment generating function of an exponential random variable?

$$\begin{split} X &\sim \mathsf{Exp}(\lambda) \\ \mathcal{M}_{\chi}(t) &= \mathsf{E}(e^{tX}) = \int_{0}^{\infty} e^{tx} \, \lambda \, e^{-\lambda x} \, dx = \int_{0}^{\infty} \lambda \, e^{(t-\lambda)x} \, dx = \frac{\lambda}{t-\lambda} \, e^{(t-\lambda)x} \Big|_{x=0}^{x=\infty} \\ &= \frac{\lambda}{t-\lambda} \left(0 - 1 \right) = \frac{\lambda}{\lambda-t} & \text{Note that} \\ &\lim_{x \to \infty} \frac{1}{t-\lambda} \, e^{(t-\lambda)x} = 0 \\ \mathcal{M}_{\chi}(t) &= \frac{\lambda}{\lambda-t} & \text{for } t < \lambda & \text{if } t-\lambda < 0. \end{split}$$

4. Let $X \sim \text{Exp}(1)$. Find a formula for $E(X^n)$ for positive integers *n*.

$$M_{x}(t) = \frac{1}{(1-t)^{-1}} = (1-t)^{-1}$$

$$M_{x}'(t) = (1-t)^{-2} \qquad E(X) = 1$$

$$M_{x}''(t) = 2(1-t)^{-3} \qquad E(X^{2}) = 2 \qquad E(X^{n}) = n!$$

$$M_{x}''(t) = 6(1-t)^{-4} \qquad E(X^{3}) = 6$$

These moments could also be computed by integrals:

$$E(X^{n}) = \int_{0}^{\infty} x^{n} e^{-x} dx$$

The integral above is related to the gamma function:

$$T(\alpha) = \int_{0}^{\infty} x^{\alpha-1} e^{-x} dx$$

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BONUS: For a positive continuous random variable *X* with pdf *f* and cdf *F*, the *hazard rate* (or *failure rate*) is defined by $h(t) = \frac{f(t)}{1-F(t)}$.

(a) Interpret the hazard rate as a conditional probability? *Hint*: $P(t < X < t + \Delta t) \approx f(t)\Delta t$

$$P(t < X < t + st | X > t) = \frac{P(t < X < t + st)}{P(X > t)} \approx \frac{f(t) st}{1 - F(t)}$$

If X is the lifetime of some item, then h(t) is the rate of failure
for items at age t.

(b) Compute the hazard rate for $X \sim \text{Exp}(\lambda)$?

If
$$X \sim E \times p(\lambda)$$
, then $h(t) = \frac{\lambda e^{-\lambda t}}{e^{-\lambda t}} = \lambda$.

Thus, the failure rate of an exponential random variable is constant. This is a consequence of the memoryless property.