- 4. Suppose that  $X_1$  and  $X_2$  are iid Unif[0,1]. Let  $Y_1 = X_1 + X_2$  and  $Y_2 = X_1 X_2$ .
- (a) Find the region of possible values of the pair  $(Y_1, Y_2)$ .



(b) Find the inverse transformation functions  $v_1$  and  $v_2$  such that  $X_1 = v_1(Y_1, Y_2)$  and  $X_2 = v_2(Y_1, Y_2)$ .

Invert the transformation matrix:  $\begin{bmatrix} 1 & 1 \\ 1 & -1 \end{bmatrix}^{-1} = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{bmatrix}$ So  $\begin{bmatrix} X_{i} \\ X_{2} \end{bmatrix} = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{bmatrix} \begin{bmatrix} Y_{i} \\ Y_{2} \end{bmatrix}$ , which means  $X_{i} = v_{i} (Y_{i}, Y_{2}) = \frac{Y_{i} + Y_{2}}{2} \quad \text{and} \quad X_{2} = v_{2} (Y_{i}, Y_{2}) = \frac{Y_{i} - Y_{2}}{2}$ 

(c) Use the transformation theorem to find the joint pdf of  $Y_1$  and  $Y_2$ .

Jacobian matrix: 
$$M = \begin{bmatrix} \frac{\partial u_{i}}{\partial y_{i}} & \frac{\partial u_{i}}{\partial y_{2}} \\ \frac{\partial v_{z}}{\partial y_{1}} & \frac{\partial v_{z}}{\partial y_{2}} \end{bmatrix} = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & -\frac{1}{2} \end{bmatrix}$$
So: 
$$det(M) = \frac{1}{2}(-\frac{1}{2}) - \frac{1}{2}(\frac{1}{2}) = -\frac{1}{2}$$
The joint density of Y<sub>1</sub> and Y<sub>2</sub> is:  

$$g(y_{1}, y_{2}) = f(v_{1}(y_{1}, y_{2}), v_{2}(y_{1}, y_{2})) \cdot |det(M)| = 1 \cdot |\frac{1}{2}| = \frac{1}{2}$$
on the region found in part (a).

5. Let (X, Y) be a random point in the plane, where X and Y are independent standard normal random variables. Let  $(R, \Theta)$  be the polar coordinates of (X, Y). Find the joint density of R and  $\Theta$ . Then find the marginal densities of R and  $\Theta$ . What is the probability that the point (X, Y) lies in a circle of radius 1 centered at the origin?



**Region:** X and Y may be any real numbers, so  $R \ge 0$  and  $0 \le \Theta \le 2\pi$ .

Joint Density: N(0,1) has pdf 
$$\frac{1}{\sqrt{2\pi}}e^{-x^2/2}$$
  
X and Y are independent, so  $f(x, y) = \frac{1}{\sqrt{2\pi}}e^{-x^2/2} \cdot \frac{1}{\sqrt{2\pi}}e^{-\gamma^2/2}$   
 $f(x, y) = \frac{1}{2\pi}e^{-(x^2+y^2)/2}$ 

From the video for today, the transformation theorem gives:

$$g(r, \theta) = f(r \cdot \cos \theta, r \cdot \sin \theta) \cdot r$$
  
Thus: 
$$g(r, \theta) = \frac{1}{2\pi} e^{-(r^2 \cos^2 \theta + r^2 \sin^2 \theta)/2} \cdot r$$

 $g(r, \theta) = \frac{r}{2\pi} e^{-r^2/2} \quad \text{for } r \ge 0, \ 0 \le \theta < 2\pi$ 

Inside circle:

$$P(R < 1) = \int_{0}^{1} r e^{-r^{2}/2} dr = -e^{-r^{2}/2} \Big|_{r=0}^{r=1} = 1 - e^{-r^{2}/2} \approx 0.393$$

1. Let  $X_1$  and  $X_2$  be iid  $\operatorname{Exp}\left(\frac{1}{10}\right)$ . (a) What is the pdf of  $Y_1 = \min(X_1, X_2)$ ?

$$g_{1}(\gamma) = n \left[1 - F(\gamma)\right]^{n-1} f(\gamma) = 2 \left[1 - (1 - e^{-\gamma/10})\right]^{1} \left(\frac{1}{10} e^{-\gamma/10}\right) = 2 \left[e^{-\gamma/10}\right] \left(\frac{1}{10} e^{-\gamma/10}\right) = \frac{1}{5} e^{-\gamma/10} f_{10} + \gamma > 0$$

$$Y_{1} \sim E_{XP} \left(\lambda = \frac{1}{5}\right)^{1/10}$$

(b) What is the expected value of  $Y_1$ ?

(c) What is the pdf of  $Y_2 = \max(X_1, X_2)$ ? What is  $E(Y_2)$ ?

 $g_{2}(y) = n \left[F(y)\right]^{n-1} f(y) = 2 \left[1 - e^{-\gamma/0}\right]^{1} \left(\frac{1}{10} e^{-\gamma/0}\right) = \frac{2}{10} \left(e^{-\gamma/0} - e^{-\gamma/5}\right) = \frac{1}{5} \left(e^{-\gamma/0} - e^{-\gamma/5}\right) \quad \text{for } y > 0$  $E(Y_{2}) = \int_{0}^{\infty} y \cdot \frac{1}{5} \left(e^{-\gamma/0} - e^{-\gamma/5}\right) dy = 15$ 



2. Let  $X_1, X_2, X_3$  be iid  $\text{Exp}\left(\frac{1}{10}\right)$ . What is the expected value of the sample median?

Sample median is  $Y_2$  (n=3, i=2)  $g_2(y) = \frac{3!}{1! 1!} \left[ 1 - e^{-\gamma/_0} \right]^1 \left[ e^{-\gamma/_0} \right]^1 (\frac{1}{10} e^{-\gamma/_0}) = \frac{6}{10} \left[ 1 - e^{-\gamma/_0} \right] e^{-\gamma/_5} = \frac{3}{5} \left( e^{-\gamma/_5} - e^{-3\gamma/_0} \right) \quad \text{for } \gamma > 0$  $E(Y_2) = \int_0^\infty \gamma \cdot \frac{3}{5} \left( e^{-\gamma/_5} - e^{-3\gamma/_0} \right) d\gamma = \frac{25}{3}$ 

3. Let  $X_1, X_2, X_3$  be iid Unif[0,1]. What is the probability that the sample median is between  $\frac{1}{4}$  and  $\frac{3}{4}$ ? n=3,  $U_{nif}[0,1]$  has pdf f(x) = 1, cdf F(x) = x, for  $0 \le x \le 1$ density of the sample median:

$$g_{2}(\gamma) = \frac{3!}{(2-i)!(3-2)!} \quad \gamma (1-\gamma) = 6\gamma (1-\gamma) \quad \text{for } 0 \leq \gamma \leq 1$$
Thus, 
$$P(\frac{1}{4} < Y_{2} < \frac{3}{4}) = \int_{\frac{1}{4}}^{\frac{3}{4}} (6\gamma - 6\gamma^{2}) d\gamma = \left[3\gamma^{2} - 2\gamma^{3}\right]_{\frac{1}{4}}^{\frac{3}{4}} = \frac{11}{16}$$

4. Let *n* be a positive odd integer and let  $X_1, X_2, ..., X_n$  be iid Unif[0,1]. What is the smallest *n* such that the sample median is between 0.4 and 0.6 with probability greater than  $\frac{1}{2}$ ?

median is 
$$Y_i$$
 with  $i = \frac{n+1}{2}$   
density:  $g_i(y) = \frac{n!}{(i-1)!(n-i)!} y^{i-1} (1-y)^{n-i} = \frac{n!}{(\frac{n-1}{2})!^2} y^{\frac{n-1}{2}} (1-y)^{\frac{n-1}{2}}$   
Use Mathematica to compute  $\int_{0.4}^{0.6} g_i(y) dy$  for various n.  
The smallest odd n such that  $\int_{a_y}^{0.6} g_i(y) dy > \frac{1}{2}$  is  $n = 11$ .

5. Let  $X_1, ..., X_8$  be iid Unif[0,1].

(a) Make a plot of the pdfs of all eight order statistics.

(b) What are the expected values of all eight order statistics?

 $E(Y_i) = \frac{i}{q}$